



MET3220C & MET6480 Computational Statistics



Lecture 8

Parametric Probability Distributions

Continuous Distributions

Key Point: ALWAYS LOOK AT THE DATA!!!!
DOES THE DATA REALLY FIT THE DISTRIBUTION?

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Parametric Probability Distributions 1

Fitting Parameters for Continuous Distributions

Distribution	E[x]	Var[x]
Gaussian	μ	σ^2
Log-normal	$\exp[\mu + \sigma^2/2]$	$(\exp[\sigma^2] - 1) \exp[2\mu + \sigma^2]$
Gamma	$\alpha\beta$	$\alpha\beta^2$
Exponential	β	β^2
Chi-squared	v	$2v$
Pearson III	$\zeta + \alpha\beta$	$\alpha\beta^2$
Beta	$p / (p + q)$	$(pq) / ((p + q)^2(p + q + 1))$
GEV	$\zeta - \beta[1 - \Gamma(1 - \kappa)] / \kappa$	$\beta^2[\Gamma(1 - 2\kappa) - \Gamma^2(1 - \kappa)] / \kappa^2$
Gumbel	$\zeta + \gamma\beta$	$\beta \pi / \sqrt{6}$
Weibull	$\beta \Gamma(1 + 1 / \alpha)$	$\beta^2[\Gamma(1 + 2 / \alpha) - \Gamma^2(1 - \kappa)] / \kappa^2$
Mixed Exponential	$w\beta_1 + (1 - w)\beta_2$	$w\beta_1^2 + (1 - w)\beta_2^2 + w(1 - w)(\beta_1 - \beta_2)^2$

μ = mean, σ = standard deviation

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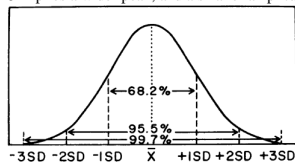
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Gaussian Distribution: The Formula

- A normal distribution is described by two parameters: a mean (μ) and a standard deviation (σ).
- A Gaussian distribution (not a pdf) would also have an amplitude.

$$pdf = f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right], -\infty < x < \infty$$

- Think about how the standard deviation influences the shape of $f(x)$.
 - Larger σ implies a wider peak, and a smaller amplitude.



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Graphic from <http://homepage.univie.ac.at/Franz.Vesely/cp0102/dk/img579.png>
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Continuous Distributions

- Continuous distributions have probabilities for any value(s) within a parameter space.
 - For example, a univariate distribution has probabilities for upper and lower bounds, as well as all values between these bounds.
 - This limits could be $\pm\infty$.
- The probability distribution function $f(x)$ is such that $\int f(x)dx = 1$.
 - Probability distribution (or density) function is abbreviated as PDF.
- Note that the probability of an event occurring is the area under the PDF, bounded by the limiting conditions on the event.
- These last two points should make it clear that $f(x) = \partial\Pr(x)/\partial x$.
 - This equation is easily written in terms of cumulative probability CDF, $C\{X \leq x\}$, because $\partial\Pr(x)/\partial x = \partial C\{X \leq x\}/\partial x$
 - If we can calculate a CDF, then we can easily randomly generate a distribution that matches the CDF and corresponding PDF.
 - Particularly so if we can determine $X(C)$ from $C(X)$.

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Gaussian Distribution

- A Gaussian distribution (bell curve) is relatively common, particularly when describing differences.
 - If a Gaussian distribution is normalized, meaning the area under the curve is equal to unity (one), then this special case of the Gaussian distribution is sometimes called a normal distribution.
 - Definitions do vary: Wilks defines the Gaussian distribution as I have defined a normal distribution.
- Estimates of a sum (or mean) will have a Gaussian distribution if the samples are (1) independent, and (2) of sufficient number.
 - The above statement is the **central limit theorem**.
 - The sufficient number is small if the population from which the samples are taken (and the sum calculated) has a near Gaussian distribution. It is larger (>100) for highly non-Gaussian PDFs.

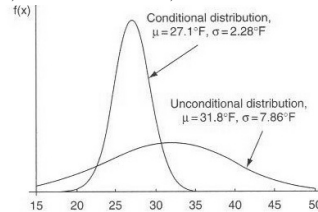
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Distributions For Conditional Probabilities

- The pdf for a conditional probability can have a very different shape than the unconditional probability.
- For example, consider the pdf for January daily maximum temperatures at Canandaigua: mean = 31.8°F, $\sigma = 7.86^\circ\text{F}$.
- If the data set is restricted to those days when the temperature at Ithaca was 25°F, then the mean is 27.1°F, and $\sigma = 2.28^\circ\text{F}$



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CDF For a Gaussian Distribution

- The technique for determining a CDF is often the integration of the corresponding pdf.

$$CDF(x) = \int_0^x pdf(x') dx'$$
- However, the Gaussian function is non-integratable.
- One approach to solving this problem is a lookup table.
 - Table B.1 in Wilks' book shows the probabilities in terms of z values: $z = (x - \mu) / \sigma$
 - z scores are numbers of standard deviations above (positive values) or below (negative values) the mean.
- The lookup table shows $Pr\{Z \leq z\}$
- Note that the Gaussian function is symmetric.
 - Therefore $Pr\{Z \leq z\} = 1 - Pr\{Z \geq -z\}$

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Approximating the Gaussian CDF

- When a good approximation is sufficient, there is a relatively simple function that can be used as an approximate CDF, $\Phi(z)$.

$$\Phi(z) = \frac{1}{2} \left[1 \pm \sqrt{1 - \exp\left(\frac{-2z^2}{\pi}\right)} \right]$$
 - Where the positive root is used for $z > 0$, and the negative root for $z < 0$
 - Where z is the number of standard deviations from the mean.
- The maximum errors (in probability) using this approximation are about 0.003 when $z = \pm 1.65$.
- This can be inverted to solve for z as a function of the value of the CDF.

$$z = \left[-\frac{\pi}{2} \ln \left[1 - [2\Phi(z) - 1]^2 \right] \right]^{1/2}$$

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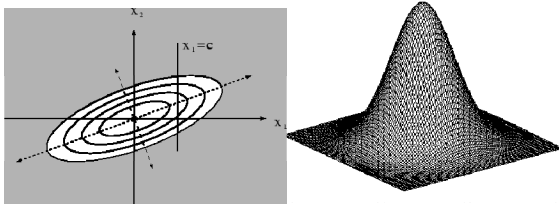


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Two Dimensional Gaussian Distributions

- Two dimensional Gaussian PDFs are also common, particularly when showing differences in two spatial dimensions.

$$pdf = f(x) = \frac{1}{\sigma_x \sigma_y \sqrt{2\pi}} \exp \left[-\frac{1}{2} \left(\frac{(x - \mu_x)^2}{\sigma_x^2} + \frac{(y - \mu_y)^2}{\sigma_y^2} \right) \right], -\infty < x < \infty$$



Graphic from <http://www.westgard.com/lesson041.htm>
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Graphic from <http://www.sia.uq.edu.au/physics/light/fred.gif>
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Log-Normal Distributions

- There are many occurrences of distributions that have
 - (1) only positive values, and
 - (2) peak is displaced to the left.
- Some of these distributions are log-normal distributions.
 - A transformation of variables is used: $y = \ln(x)$

$$pdf = f(x) = \frac{1}{x \sigma_y \sqrt{2\pi}} \exp \left[-\frac{(\ln(x) - \mu_y)^2}{2\sigma_y^2} \right], -\infty < y < \infty, y = \ln(x)$$

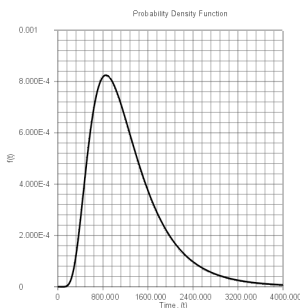
- Where μ_y and σ_y are the mean and standard deviation of the transformed variable y.
- The mean of x is $\exp[\mu + \sigma^2/2]$, and the standard deviation of x is $(\exp[\sigma^2] - 1) \exp[\mu + \sigma^2]$.
- Where μ and σ are the mean and standard deviation of the transformed variable y.

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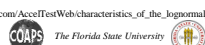
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Log-Normal Distribution Example



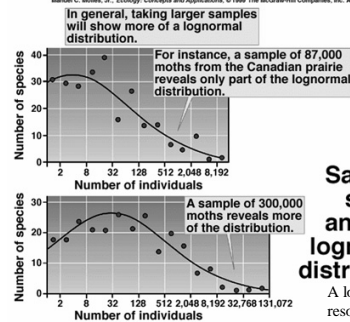
- Features:
 - (1) only positive values, and
 - (2) peak displaced to the left.
- If the x-axis was plotted in log coordinates, then the distribution would appear to be Gaussian.

Graphic from http://www.weibull.com/AceTestWeb/characteristics_of_the_lognormal_distribution.htm
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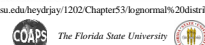
Log-Normal Distribution Example



Sample size and the lognormal distribution.

A lot more data helps resolve extremes

Graphic from <http://www.biology.lsu.edu/hedyjaya/1202/Chapter53/lognormal%20distribution.jpg>
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Gamma Distributions

- Gamma distributions are asymmetric, and skewed to the right (meaning the peak is to the left of the mean).
- They are well suited to describe variables that have a peak close to a limit.
 - For example, wind speed or precipitation.
- There are several different (but equivalent) forms of the gamma distribution. Each has two fitting parameters
- The fitting parameters are a shape parameter α , and a scaling parameter β .
 - Alternatively, it can be written with an inverse scale factor.

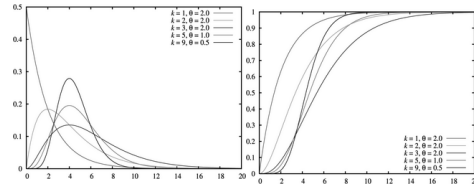
$$f(x) = \frac{(x/\beta)^{\alpha-1} \exp(-x/\beta)}{\beta \Gamma(\alpha)}, \text{ for } x, \alpha, \beta > 0$$

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Gamma Distribution



- The above examples use k and the shape parameter, and θ as the scale parameter.
- The left plot is the PDF, and the right plot is a CDF
- For a constant scale parameter, a smaller shape parameter will result in the peak being shifted further to the left

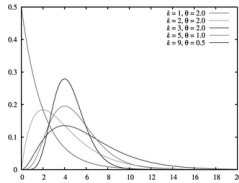
$$f(x; k, \theta) = x^{k-1} \frac{e^{-x/\theta}}{\theta^k \Gamma(k)} \text{ for } x > 0$$

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Gamma Distribution Parameters



$$f(x; k, \theta) = x^{k-1} \frac{e^{-x/\theta}}{\theta^k \Gamma(k)} \text{ for } x > 0$$

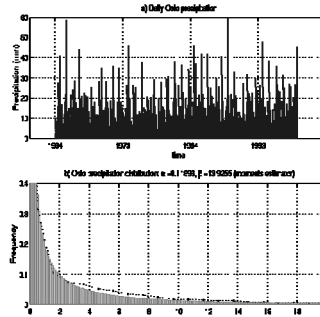
- For a shape parameter $k=1$, the equation simplifies greatly to an exponential distribution.
 - The y-intercept is $1/\theta$.
- For a shape parameter $k > 1$, the y-intercept is zero.
 - Larger values of k result in less skewness, and shift the peak to the right.
 - For $k > 50$ or 100 , the distribution is approximately Gaussian.

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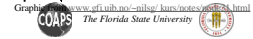
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Gamma Distribution Example



- Time series of daily precipitation at Oslo (top)
- The distribution function for daily precipitation in Oslo between 1883 and 1964 (bottom), with the dashed line showing the distribution for the above time period.

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Estimating the Gamma Distributions Scale Parameter

- We want to determine the fitting parameters α and β .
- We can solve for these in terms of the mean and the standard deviation of the gamma function.

$$\bar{x} = \alpha \beta$$

$$\sigma = \alpha \beta^2$$

- We can solve these equations for the fitting parameters:

$$\alpha = \bar{x}^2 / \sigma^2$$

$$\beta = \sigma^2 / \bar{x}$$

- What could go wrong with this approach?
 - Good for (shape parameter) $\alpha > 10$
 - Poor estimates of moments lead to problems for smaller α .

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More Robust Estimates of Fitting Parameters

- Two better methods are based on *maximum likelihood estimators*.
 - This concept will be explained in later lectures
- Both approach use the same 1st calculation

$$D = \ln(\bar{x}) - \frac{1}{n} \sum_{i=1}^n \ln(x_i)$$

- The Thom estimators are

$$\alpha = \frac{1 + \sqrt{1 + 4D/3}}{4D} \text{ and } \beta = \bar{x} / \alpha$$

- The other method (Greenwood and Durand, *Technometrics*, 1960)

$$\alpha = \frac{0.5000876 + 0.1648852 D - 0.0544274 D^2}{D}, \quad 0 \leq D \leq 0.5772$$

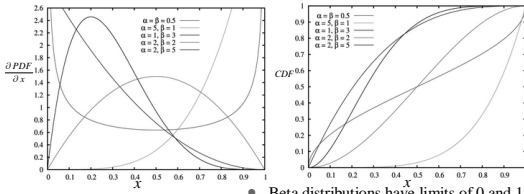
$$\alpha = \frac{8.898919 + 9.059950 D + 0.9775373 D^2}{17.19728 D + 11.968477 D^2 + D^3}, \quad 0.5772 \leq D \leq 17.0$$

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Beta Distributions



- Beta distributions have limits of 0 and 1.
 - Applications: RH & cloud cover
- They have two tuning parameters: α, β .
- The B term normalizes the PDF.
- If $\alpha = \beta$, the distribution is symmetric.
- If α and β are exchanged, the $f(x)$ is mirrored around $x = 0.5$.

$$f(x; \alpha, \beta) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

$$f(x; \alpha, \beta) = \frac{1}{\int_0^1 u^{\alpha-1} (1-u)^{\beta-1} du} x^{\alpha-1} (1-x)^{\beta-1}$$

$$= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

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Extreme Value Distributions

- Extreme value distributions usually apply to a small fraction of the events: the extreme events.
 - E.g., floods at a specific location
- The fraction can be artificially increased by using only extreme values in the distribution.
 - E.g., the annual maximum of daily precipitation totals (at a specific location).
- The Generalized Extreme Value (GEV) Distribution is

$$f(x) = \frac{1}{\beta} \left[1 + \frac{\kappa(x-\zeta)}{\beta} \right]^{-1/\kappa} \exp \left\{ - \left[\frac{\kappa(x-\zeta)}{\beta} \right]^{-1/\kappa} \right\}$$

- Where ζ is a location or shift parameter, β is a scale parameter, and κ is a shape parameter.

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CDF of a GEV Distribution

- The GEV equation can be integrated, resulting in an analytical CDF.

$$CDF(x) = \exp \left\{ - \left[1 + \frac{\kappa(x-\zeta)}{\beta} \right]^{-1/\kappa} \right\}$$
- The CDF can be inverted (solved for x as a function of $CDF(x)$).

$$CDF^{-1}(p) = x = \zeta + \frac{\beta}{\kappa} \left[\left(-\ln(p) \right)^{\kappa} - 1 \right]$$
- Given the fitting parameters, we can determine the extreme value as a function of the probability of that extreme (or greater) occurring.
 - We don't expect the distribution to work for likely occurrences.
 - However, as p becomes smaller, the distribution can be quite realistic.
 - Note that as $p \rightarrow 0$, that $\ln(p) \rightarrow -\infty$, resulting in rather large x .
- There are three special cases of the GEV Distribution. The two that we will examine are the Gumbel distribution and the Weibull distribution.

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Gumbel Distribution

- Typically used to determine the average time between extreme events of the same magnitude or greater.
- The Gumbel distribution is the limit of the GEV distribution, where $\kappa \rightarrow 0$.

$$f(x) = \frac{1}{\beta} \exp \left\{ - \exp \left[\frac{(x-\zeta)}{\beta} \right] - \frac{(x-\zeta)}{\beta} \right\}$$

$$CDF(x) = \exp \left\{ - \exp \left[\frac{(x-\zeta)}{\beta} \right] \right\}$$
- The fitting parameter can be estimated through a method of moments.
 - $\beta = \sigma \sqrt{6} / \pi$
 - $\zeta = \bar{x} - \gamma \beta$
 - Where $\gamma = 0.57721\dots$ is Euler's constant.

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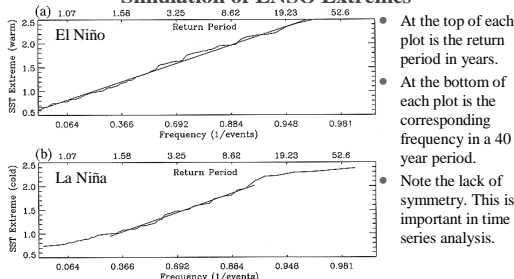


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Gumbel Distribution Example: Simulation of ENSO Extremes



- Statistical mumbo-jumbo was used to generate 40 years of a sea surface temperature based ENSO index.

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Graphics from Caron and O'Brien, Mon. Wea. Rev. 126:2809-2821.



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Return Period

- The return period is **average** time between events of a certain magnitude or greater.
- Note that the return period is an average. Three 100-year flood events have been known to happen in within 5 years.
- Suggesting that there might be year-to-year memory of ground water conditions.
- The return period R for an event of magnitude x or greater is $R(x) = 1 / \{\omega[1 - CDF(x)]\}$
 - Where ω is the sampling interval.

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Weibull Distribution

- Weibull distributions are the limit of the GEV distribution where $\kappa < 0$.
- They have the distribution

$$f(x) = \left(\frac{\alpha}{\beta}\right) \left(\frac{x}{\beta}\right)^{\alpha-1} \exp\left[-\left(\frac{x}{\beta}\right)^\alpha\right]$$

$$CDF(x) = 1 - \exp\left[-\left(\frac{x}{\beta}\right)^\alpha\right]$$

- The method of moments does not work for determining the fitting parameters. The gamma functions awkward.

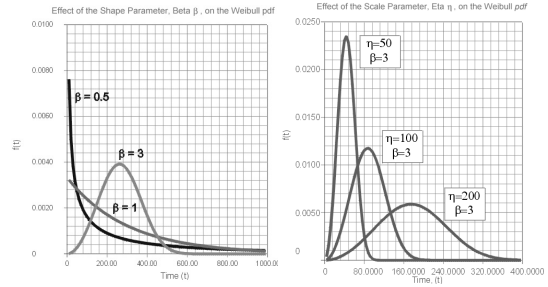
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Weibull Distribution Examples

Graphics from www.weibull.com/basics/parameters.htm



- In this example the shape parameter is β (our α), and the scale parameter is η (our β).

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Mixtures of Distributions

- For mildly complex physical situations, there is no reason that one type of distribution should fit the data.
- If there are several processes contributing to the physics (e.g., processes for generating rain), then it might be necessary to use a weighted average of several distributions.
- Example: $wt1 * (\text{Gaussian Distribution 1}) + wt2 * (\text{Gaussian Distribution 2}) + (1 - wt1 - wt2) * \text{Weibull Distribution}$
 - Where $0 < wt1 < 1$, $0 < wt2 < 1$, and $0 < wt1 + wt2 < 1$

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